

Contrasting patterns of environmental fluctuation contribute to divergent life histories among amphibian populations

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1	Running head:
2	HABITAT PREDICTABILIY AND LIFE HISTORY TACTICS
3	Title:
4	Contrasting patterns of environmental fluctuation contribute to divergent life histories among
5	amphibian populations
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Abstract. Because it modulates the fitness returns of possible options of energy expenditure at 22 23 each ontogenetic stage, environmental stochasticity is usually considered a selective force in 24 driving or constraining possible life histories. Divergent regimes of environmental fluctuation 25 experienced by populations are expected to generate differences in the resource allocation 26 schedule between survival and reproductive effort and outputs. To our knowledge, no study 27 has previously examined how different regimes of stochastic variation in environmental conditions could result in changes in both the temporal variation and mean of demographic 28 29 parameters, which could then lead to intra-specific variation along the slow-fast continuum of 30 life history tactics. To investigate these issue, we used capture-recapture data collected on 31 five populations of a long-lived amphibian (Bombina variegata) experiencing two distinct levels of stochastic environmental variation: (i) constant availability of breeding sites in space 32 and time ('predictable environment'), and (ii) variable spatio-temporal availability of breeding 33 34 sites ('unpredictable environment'). We found that female breeding propensity varied more from year to year in unpredictable than in predictable environments. Although females in 35 36 unpredictable environments produced on average more viable offspring per year, offspring production was more variable between years. Survival at each ontogenetic stage was slightly 37 38 lower and varied significantly more from year to year in unpredictable environments. Taken together, these results confirm that increased environmental stochasticity can modify the 39 40 resource allocation schedule between survival and reproductive effort and outputs and may 41 lead to intra-specific variation along the slow-fast continuum of life history tactics.

42

43 *Key words: life history; slow–fast continuum; environmental predictability; demography;*

44 multievent capture-recapture models; amphibian; Bombina variegata

45 INTRODUCTION

46 Identifying the rules that shape life history evolution is a crucial challenge for evolutionary 47 biology. According to life history theory, ontogenesis involves a sequence of developmental 48 events that are ordered to optimize survival and reproductive outputs (Stearns 1992). Over a 49 lifetime, energy and trophic resources are allocated between growth, somatic maintenance and reproductive effort (Pianka 1976). Growth and somatic maintenance include investments that 50 contribute to longevity, whereas reproductive effort includes investments in offspring 51 52 production and assistance that contribute to fecundity. Because it modulates the fitness returns 53 of possible options of energy expenditure at each ontogenetic stage, stochastic environmental 54 variation is usually considered a selective force in driving or constraining possible life 55 histories (Levins 1968, Schaffer 1974, Tuljapurkar et al. 2009).

56 When environmental fluctuation generates strong variation in adult mortality, natural selection favors a short lifespan, early maturity and large reproductive outputs (characteristics 57 of 'fast' life histories along a 'slow-fast continuum'; e.g. Kraus et al. 2005). In contrast, when 58 59 environmental fluctuation affects birth rate and juvenile mortality, it contributes to the 60 selection for a longer lifespan, later maturity and smaller reproductive outputs (characteristics of 'slow' life histories; e.g. Morris et al. 2008). For species with long lifespans, selective 61 62 forces drive the evolution of traits that favor a canalization process (sensu Gaillard and 63 Yoccoz 2003) of adult survival since this demographic parameter has the highest impact on 64 fitness. Such a reduction of variation in adult survival may initially be achieved by trade-offs 65 between vital rates (Stearns 1992, Shaw and Levin 2011). In particular, adults may modulate 66 their yearly breeding investment or may temporarily cease to breed at any age to increase longevity (Warner 1998, Rideout et al. 2005, Ruf et al. 2006). Reduced variation in adult 67 68 survival may also result from behavioral or physiological processes that buffer the impact of environmental fluctuation. For instance, this can be seen in adaptive behavioral mechanisms 69

such as the wing-fanning and shade-seeking of endotherms that allow individuals to maintain
their body temperature in increasing ambient temperatures, which may drastically reduce
adult survival variation due to temperature fluctuation (Welbergen et al. 2008).

73 By constraining the availability and quality of resources that could be used by an 74 iteroparous organism over its lifetime, environmental fluctuation also regulates the spread of 75 reproductive effort over time (Tuljapurkar et al. 2009). An organism's ability to reproduce strongly depends on a combination of environmental and physiological factors (McNamara 76 77 and Houston 1996), which include its fat reserves, foraging skills, parasite load and the 78 responsiveness of its immune system. Given these constraints, at each breeding opportunity 79 an iteroparous organism has to make a decision whether or not to breed, and if it does breed, it must modulate its breeding investment in accordance with its lifetime fitness gain (Reznick 80 and Yang 1993, Muths et al. 2010, Souchay et al. 2014). Variation in food availability due to 81 82 environmental fluctuation (e.g. yearly weather variation) may lead to insufficient energy accumulation for initiating oocyte development in females, which therefore skip reproduction 83 84 (Jenouvrier et al. 2003). Moreover, environmental fluctuation also regulates the availability and quality of resources required by offspring during their development, which may lead 85 86 organisms to modulate breeding investment accordingly. Breeding females may respond to fluctuation in food availability by delaying oocyte production, which increases their fitness as 87 this produces offspring with higher fat content, which can better survive and grow in low-food 88 89 environments (Reznick and Yang 1993).

As is predicted at inter-specific level (Stearns 1992), at intra-specific level, divergent regimes of environmental conditions experienced among populations are expected to generate differences in the resource allocation schedule between survival and reproductive effort and outputs (Van Tienderen 2000). For instance, a population of *Capreolus capreolus* submitted to the harsh environmental conditions generated by the co-occurrence of high hunting

pressure and predation by lynx exhibited a slower life history cycle than that of a population 95 96 in less harsh conditions (Nielsen et al. 2009). Conversely, an acceleration of the life history 97 cycle has been reported in Sarcophilus harrisii; the population responded to increased adult mortality due to a pathogen by increasing reproductive effort at an earlier age, implying an 98 99 abrupt transition from iteroparity to semelparity (Jones et al. 2008). These studies provide an 100 interesting overview on how contrasting environmental conditions can impact – sometimes very quickly – the trade-off between mean survival and reproductive effort and outputs, and in 101 102 turn affect the speed of the life history cycle at intra-specific level. However, no previous 103 study has yet examined how different regimes of stochastic variation in environmental 104 conditions could result in changes in both the temporal variation and mean of demographic parameters, which could then lead to intra-specific variation along the slow-fast continuum of 105 106 life history tactics.

The aim of this study was to test if contrasting levels of environmental fluctuation 107 contribute to divergent life histories among free-ranging populations of a long-lived 108 109 amphibian, the yellow-bellied toad (Bombina variegata). We conducted a study on five 110 populations of *B. variegata* including two levels of stochastic environmental variation: (i) 111 constant availability of breeding sites in space and time (called hereafter 'predictable environments' as breeding site availability was predictable from one year to another) and (ii) 112 113 stochastic spatio-temporal availability of breeding sites, which were randomly created or 114 destroyed each year by human activity (called hereafter 'unpredictable environments' as 115 breeding site availability was not predictable from year to year). In each population studied, 116 we considered the following key parameters: (i) age-dependent survival, (ii) female breeding 117 propensity, (iii) offspring productivity per female per year (a proxy for female fitness) and (iv) temporal variation of these vital rates. Using capture-recapture data collected in both 118 predictable and unpredictable environments, we specifically tested the hypothesis that 119

survival would be lower on average and more variable from year to year in unpredictable 120 121 environments, whatever the ontogenetic stage considered (juvenile, subadult or adult); 122 increased mortality was expected to result from human activity (e.g. through pollution or habitat alteration) (Mann et al. 2009, Semlitsch et al. 2009) and/or from the cost induced by 123 124 moving from an aging rut network to a more recent one (Bonte et al. 2012). We also tested the hypothesis that a female could modify its schedule of resource allocation to reproduction 125 126 according to the degree of environmental fluctuation. In particular, we predicted that breeding 127 propensity would be more variable from year to year in unpredictable environments. In 128 parallel, we hypothesized that offspring productivity would also be more variable from year to 129 year, but would be higher on average in unpredictable environments. Accordingly, we expected that enhanced environmental stochasticity could result in an accelerated life history 130 cycle. 131 132 133 **METHODS** 134 Biological model

135

136 *Bombina variegata* is a long-lived anuran. The life expectancy ranges from 4 to 10 years in natural conditions given the estimated survival rate of 0.7–0.9 (Beshkov and Jameson 1980, 137 138 Cayuela et al. 2014, 2015b and the present study); some individuals reach an age of more than 139 20 years in the wild and almost 30 years in captivity (Smirina 1994). This species breeds in small water bodies characterized by stochastic water level variation. Its long lifespan and 140 141 iteroparity are assumed to be a response to habitat uncertainty (Joly and Morand 1994, 142 Morand and Joly 1995, Cayuela et al. 2014). Furthermore, females display risk-spreading tactics such as egg clutch fractionation over space and time (Buschmann 2002). The 143

flexibility of the ovarian functioning of bombinatorids allows the continuous production or
resorption of oocytes over the course of the breeding season (Guarino et al. 1998).

- 146In Western Europe, *Bombina variegata* populations occupy both natural (Cayuela et
- al. 2011, 2013) and artificial environments (Cayuela et al. 2014, 2015*a*). In natural
- 148 environments such as riverbanks (usually considered as ancestral habitats), these toads breed
- in rocky pools and natural depressions filled by floodwater or rainwater. The availability of
- 150 natural breeding sites depends on long-term spatio-temporal changes in waterbed
- morphology; on the scale of a toad's lifetime, natural breeding site availability is therefore
- 152 highly predictable ('predictable environment'). In contrast, *B. variegata* individuals that
- 153 occupy artificial environments, such as woodlands exploited for logging or agricultural areas,
- breed in artificial water bodies such as ruts, ditches and residual puddles created by skidders
- 155 or farming equipment. Thus the inter-annual predictability of potential breeding sites depends
- 156 on short-term spatio-temporal variation driven by logging or farming operations. In exploited
- 157 woodlands, for instance, rut networks are created during logging operations (Cayuela et al.
- 158 2015a) and are filled in at the end of the exploitation of the forest patch or in the following
- two years; on the scale of an individual's lifespan, the availability of breeding sites is
- therefore highly stochastic in artificial environments ('unpredictable environment').
- 161
- 162 *Surveyed populations and fieldwork*
- 163

As in most studies that compare intraspecific and interspecific life history strategies (e.g. Gaillard and Yoccoz 2003, Nilsen et al. 2009, Miller et al. 2011, but also see Lawson et al. 2015 and Barraquand and Yoccoz 2013), we indirectly assessed environmental stochasticity through variations in demographic rates. No specific measures of environmental predictability (e.g. pond duration) were available for the studied populations. The study was conducted on

169	five <i>B. variegata</i> populations in France (TABLE 1, see also in Appendix A for a map and
170	details concerning environmental characteristics): three in unpredictable environments and
171	two in predictable environments. The three 'unpredictable environment' populations
172	(hereafter POP1, POP2 and POP3) were located in the Ardennes, Meuse and Charente
173	departments respectively. In these areas the species reproduces in networks of water bodies
174	resulting from logging and farming operations (e.g. ruts, ditches and residual puddles); these
175	are distributed in a landscape matrix consisting of managed woodlands and agricultural areas
176	(Cayuela et al. 2015a). The two 'predictable environment' populations were located in two
177	distinct watersheds in the Ardèche department. Here, the species breeds in networks of rocky
178	pools and natural depressions filled by floodwater or rainwater in the vicinity of river
179	channels (Cayuela et al. 2011, 2013). In these areas, the surrounding landscape mainly
180	consists of non-managed woodlands.
181	The five populations were monitored using capture-recapture methods during survey
182	periods varying from 5 to 10 years between 2000 and 2014 (see TABLE 1). The capture-
183	recapture (CR) survey design has previously been described in published studies (POP1,
184	Pichenot 2008; POP2, Cayuela et al. 2014); a similar design was used for POP3, POP4 and
185	POP5 (details concerning the survey design are provided in Appendix A, TABLE A1 - A6). It
186	consisted of capturing the toads by hand or with a dip net during daylight hours (09:00-
187	19:00). Capture sessions were carried out during the breeding season from late May to July.
188	The entire study area was carefully checked every year in order to census all available water
189	bodies. The yellow-bellied toads were captured in all distinct water-body networks (ranging
190	from 3 networks in POP5 to 80 networks in POP2), which were made up of between 2 water
191	bodies (POP2, POP4 and POP5) to more than 100 water bodies (e.g. in POP4 and POP5). At
192	each capture session, we sampled all water bodies in all networks. The time required to
193	sample the entire study area in a single session varied from 1 (POP3) to 15 days (POP2). We

194	considered three age classes: juveniles (captured after their first overwintering), subadults
195	(captured after their second overwintering) and sexually mature adults (captured after at least
196	three overwinterings). Although one study has considered that individuals can be sexually
197	mature at the age of 2 (Barandun et al. 1997), four others consider the sexual maturity of this
198	species to be the age of 3 (Kapfberger 1982, 1984, Rafinska 1991, Plytycz and Bigaj 1993). In
199	any case, even if gonads may already be producing viable gametes at the age of 2, we
200	assumed that individuals would be only marginally involved in breeding before the age of 3
201	because male reproductive success and female fecundity often increases with size/age in
202	amphibians (Wells 1977, Arak 1983, Woolbright 1983, Wells 2010). Therefore, we have
203	assumed that yellow-bellied toads become sexually mature at the age of 3. The size (snout-
204	vent length) of juveniles ranged from 22–29 mm, the size of subadults ranged from 30–34 m
205	and individuals become sexually mature with a mean body length of 35 mm in males and 36
206	mm in females. We identified gender on the basis of forearm size and the presence of nuptial
207	pads in males (Kyriakopoulou-Sklavounou et al. 2012). We identified each individual by the
208	specific pattern of black and yellow mottles on its belly, which were recorded by photographs.
209	To minimize misidentification errors, multiple comparisons of individual patterns were
210	performed using the pattern-matching software Extract Compare (Hiby and Lovell 1990).

211

212 *Multievent model design*

213

The CR data was modeled using a re-parameterized version of the robust design proposed by Cayuela et al. (2014). The model follows the typical structure of 'standard' robust design (Nichols et al. 1994, Kendall et al. 1997, 2012), i.e. made up of two nested levels of capture occasions. 'Secondary sessions' encompass several field sessions performed over the same year. Sampling sessions were also conducted over successive years: 'primary periods'

correspond to yearly sessions. During a primary period (i.e. between secondary sessions), we
assumed that the state of an individual was fixed and transitions between states were thus
allowed between primary periods only.

The model was conditional on the occasion of first capture and an 'unborn individual' 222 state was included, which allowed us to model the probability of being recruited as a juvenile 223 between two primary periods. After being recruited, individuals could successively transition 224 between three ontogenetic stages, i.e. juvenile, subadult and adult, given they survived. As is 225 226 typical in a capture-recapture framework, survival estimates include both death and 227 permanent emigration, i.e. apparent survival: see Schmidt et al. (2007) for a discussion on this 228 point in amphibians. We assumed that sexually mature Bombina variegata individuals could breed or skip breeding each year (see Cayuela et al. 2014). Since individuals were captured in 229 aquatic breeding sites only, non-breeding toads were considered as not available for capture 230 (Church et al. 2007, Muths et al. 2010). Skipping breeding in a given year was thus equivalent 231 to becoming a temporary emigrant and shifting to an unobservable state (Kendall and Nichols 232 233 2002). In the model, transitions between breeding/non-breeding states were assumed to obey a 234 Markov chain process.

235 We specified our states as conditional on the occasion of first capture (TABLE 2). Because gender is not identifiable in juveniles or subadults, we used a multi-event CR 236 237 approach (Pradel 2005) to model uncertainty in gender assignment; gender was thus coded as 238 a state instead of as a group as is usual in CR data modeling. Individuals could occupy one of 239 the eleven distinct states: (UM) unborn male, (UF) unborn female, (JM) alive juvenile male, 240 (JF) alive juvenile female, (SM) alive subadult male, (SF) alive subadult female, (BM)241 breeding male, (BF) breeding female, (NBM) non-breeding adult male, (NBF) non-breeding 242 adult female, (†) dead individual. Ten state transitions were possible for the recruitment and age transition probabilities between two primary periods (FIG.1, matrix A). Given that we 243

244	examined recruitment without accounting for gender-dependency, α^{JM}/α^{JF} were constrained
245	to be equal (the parameter is hereafter quoted α). Moreover, age transition probabilities
246	$\beta^{JM}/\beta^{JF}/\beta^{SM}/\beta^{SF}$ were fixed at 1 between secondary sessions and at 0 between primary
247	periods. For the survival probability between two primary periods, ten state transitions were
248	possible (FIG.1, matrix Φ). Given that the parameters ϕ^{BNM} and ϕ^{BNF} were not identifiable,
249	ϕ^{BM}/ϕ^{BNM} and ϕ^{BF}/ϕ^{BNF} were constrained to be equal in the model. Survival transition
250	probabilities were fixed at 1 between secondary sessions. For breeder/non-breeder transitions
251	between two primary periods, four state transitions were possible (FIG.1, matrix Y). The
252	transition probabilities γ^{BM}/γ^{BF} and $\gamma^{NBM}/\gamma^{NBF}$ were fixed at 1 between secondary sessions.
253	Concerning field observation, i.e. events sensu multi-event modeling formulation, six
254	observation probabilities were considered (FIG.1, matrix P). Note that as non-breeding
255	individuals were not available for recapture, their recapture probability was fixed at 0.
256	
257	Building biological scenarios
258	

The parameterization was implemented in the E-SURGE program (Choquet et al. 2009),
which provides robust tools of advanced numerical convergence and refines parameter
estimates by detecting redundant mathematical parameters.

The datasets from the five populations differ markedly in terms of the number of study years and capture sessions per year (Appendix A, TABLE A2–A6) and we did not expect any identical parameters in the different populations. Thus, the expected benefit in terms of precision or statistical power of simultaneously analyzing the five datasets would be rather small or null, while the cost in terms of computation would be high. As a result, we separately analyzed the five datasets and then compared the mean and variance of vital rates provided by the best-fitted model for each population.

269	For each population, competing models were ranked through a model-selection
270	procedure using Akaike information criteria adjusted for a small sample size (AICc) and AICc
271	relative weights (Burnham and Anderson 2002). We tested our hypotheses concerning
272	recapture and state transition probabilities from the general model [α (YEAR), (YEAR ×
273	AGE), γ (YEAR × SEX), p (YEAR + AGE + SEX)]. Variation in survival between ontogenetic
274	stages (AGE) was tested by specifying state transition differences between the three stages
275	coded as states in the model. Since we were interested in investigating female breeding
276	investment, we accounted for gender-dependent variation (SEX) in breeding probability by
277	considering state transition differences between genders coded as states in the model. Initial
278	state probability was held to be different between ontogenetic stages and gender. Furthermore,
279	recruitment probability α was held to be different among years (YEAR).
280	From the general model, we successively modeled field observation, breeding
281	probability and survival probability. We used a downward testing procedure by removing the
282	effects one by one and retained the best linear combination at each step. First, we tested
283	whether recapture probability p varied between years (YEAR), as well as between ontogenetic
284	stages (AGE) and gender in sexually mature adults (SEX). For the last hypothesis, p^{JM}/p^{JF} ,
285	p^{SM}/p^{SF} and p^{BM}/p^{BF} were held to be different. Since we aimed at investigating the yearly
286	breeding propensity of females, we then considered gender-specific differences in breeding
287	probability and tested whether breeder/non-breeder transitions γ varied between years
288	(YEAR). For this last hypothesis, the pair of transitions was held to be different γ^{BM}/γ^{BF} and
289	$\gamma^{NBM}/\gamma^{NBF}$. We then tested whether survival probability differed (1) between years
290	(YEAR) and (2) between ontogenetic stages (AGE). For the third hypothesis, ϕ^{JM}/ϕ^{JF} ,
291	ϕ^{SM}/ϕ^{SF} and ϕ^{SM}/ϕ^{SF} were constrained to be equal.
292	Because the violation of the closure assumption may bias survival estimates in robust

design models (Kendall 1999), we assessed the robustness of our analysis by removing

secondary sessions in which only a few individuals were encountered. Within a primary 294 period, the secondary session resulting in the maximum number of encountered toads has 295 296 been retained as the reference session for this year. For each primary period, we removed all 297 secondary sessions in which the number of captured toads was less than 50% of the number of toads encountered during this reference session. This led to the exclusion of between 5 and 10 298 299 sessions (detailed in Appendix B, TABLE B1) depending on the population and the year (mainly early or late secondary sessions during the breeding season). Then we fitted to these 300 301 reduced datasets the best-supported model from our model selection procedure for each 302 population. The comparisons of mean and temporal survival variance with and without the 303 removal of these secondary sessions are provided in the appendix (Appendix B, FIG. B1 and FIG. B2). Only marginal discrepancies were detected, so the potential violation of the closure 304 assumption is not an issue in our analysis. Moreover, unlike many other amphibians (Ranids, 305 306 Bufonids and Hylids), the yellow-bellied toad has a lateral line as well as complete interdigital webs (Seidel et al. 2001), which are adaptations for an extended aquatic breeding period. 307 308 Prolonged spermiation and the presence of post-vitellogenic follicles (Guarino 1998) throughout the period of breeding activity allow this species of toad to reproduce from April 309 310 to August (Di Cerbo and Biancardi 2004), which makes individuals available for capture during this relatively long period. Taken together, these factors suggest that the closure 311 assumption may hold in our study 312 313

314 Quantifying temporal variation in survival and breeding propensity

315

We built models with yearly variation in survival and breeding transition for all five populations [α (YEAR), (AGE x YEAR), γ (AGE x YEAR), p(CST)] in order to obtain annual estimates of survival rates along with their estimated variance and covariance. Using

the methods proposed in Gould and Nichols (1998), we estimated the temporal variance in 319 survival of all age classes by removing the effects of sampling variance from the estimated 320 321 total variance in survival. We used these calculated standard deviations resulting from the yearly variation in survival probability as indicators of variation in survival of a certain age 322 class over time, which was in turn a proxy for estimating the presence of canalization of the 323 given trait. We then used the method proposed by Gaillard and Yoccoz (2003) to compare the 324 325 temporal variation variance of survival rates between ontogenetic stages and populations. This 326 method compares the maximum variance of survival we can expect (from the observed mean 327 survival rate) with the observed variance in juvenile, subadult and adult survival. This 328 maximum variance is $s \times (1 - s)$, where s is the mean survival rate. Between the observed variance and the maximum variance, we then obtained the ratio of possible variation (called 329 'temporal variation ratio' hereafter, following the terminology used by Gaillard and Yoccoz, 330 331 2003) and compared this between ontogenetic stages and populations. In order to provide uncertainty measures of the temporal variation ratio of survival (and breeding) rates, we 332 333 bootstrapped in the individual capture-recapture histories. Because E-surge does not provide an automatic procedure for bootstrapping CMR-histories, we had to perform this manually. 334 335 We thus limited our approach to 50 bootstraps per population. For each bootstrap of each population, we fitted the best model we obtained from the model selection procedure and 336 applied the Gould and Nichols (1998) method. This allowed us to obtain a variance for the 50 337 338 bootstraps, from which we derived the 95% CI of the temporal variation ratio of survival rate. 339 We then used the overlap of the 95% CIs to discuss the statistical differences between the 340 variation estimates. Temporal variation ratios of breeding propensity were calculated using 341 the same procedure.

342

343 *Quantifying annual offspring production and its temporal variance*

344

345	From the same model, annual offspring production was estimated as the number of newly
346	recruited juvenile females at t per number of breeding females at $t - 1$. The number of
347	recruited juvenile females at time t (<i>NRJ</i>) and the number of annual breeding females at $t - 1$
348	(NBF) were derived (i) from the number of juveniles captured at t and breeding females
349	captured at $t - 1$, (ii) from their capture probability estimates at a given secondary session
350	(<i>CPJ</i> : the capture probability of juveniles at a secondary session of year <i>t</i> ; <i>CPBF</i> : the capture
351	probability of breeding females at a secondary session of year $t - 1$), (iii) from the best
352	supported model, and (iv) from the number of secondary sessions at $t(K)$ and $t - 1(L)$.
353	Annual offspring production was thus expressed as: $(NRJ/(1 - (1 - CP)^{K}))/(NBF/(1 - (1 - CP)^{K}))$
354	<i>CPBF</i>) ^L)). The variance and covariance of annual fecundity was estimated using the methods
355	as described for survival probability.
356	
357	RESULTS
358	
359	The model selection procedure and best-supported model for each population are provided in

360 TABLE 3. We did not detect differences in the average breeding propensity of females 361 (breeding vs. skipping reproduction during two consecutive years) between unpredictable and predictable environments (FIG. 2). However, small overlaps in the 95% CI of the temporal 362 363 variation ratios (FIG. 2) indicated that breeding propensity was more variable over time in unpredictable than in predictable environments. Furthermore, although breeding females in 364 365 unpredictable environments annually produced more offspring on average, offspring production was more variable over time (small overlaps in the 95% CI, see FIG. 3). 366 367 In the five populations, survival differed between ontogenetic stages (juvenile, subadult and adult) and increased with age (FIG. 4). At the three ontogenetic stages, survival 368

was lower in unpredictable environments. Small overlaps in the 95% CI of the temporal
variation ratios indicated that the survival of juveniles and subadults was more variable in
unpredictable environments, while large overlaps in the 95% CI suggested that temporal
variability of adult survival was similar in both environments (FIG. 4).

373

374 DISCUSSION

375

376 Comparing the population dynamics in unpredictable environments (variable spatio-temporal 377 availability of breeding sites created by human activity) and predictable environments 378 (persistent availability of natural breeding sites) reveals divergent life history tactics in B. *variegata*. While on average female breeding propensity did not differ between environments, 379 our results demonstrated that it was more variable from year to year in unpredictable than in 380 381 predictable environments. In addition, although females produced on average more viable offspring per year in unpredictable environments, offspring production was more variable 382 383 from year to year. Survival at the three ontogenetic stages was slightly lower and more variable (except in adults) from year to year in unpredictable environments. Taken together, 384 385 the differences recorded in vital rates indicate that populations in unpredictable environments respond to increased overall mortality by increasing offspring production, leading to an 386 acceleration of the life history cycle. 387

388

389 Alternative hypothesis

390

391 It is possible that alternative mechanisms – for example, differences in weather, predation or
392 competition regimes – may have caused the observed pattern in demographic rate mean and
393 variation. Weather is often considered the most important environmental driver of amphibian

vital rates (reviewed in Walls et al. 2013). Yet in our studied populations, the strongest yearly 394 395 weather variations were recorded for populations in which the spatial configuration of ponds 396 was most stable over time and in which vital rates were the least variable (Appendix A, 397 TABLE A1). This suggests that the differences of mean and variance in vital rates in our 398 findings did not result from local differences in weather variability. Another possible factor 399 could be local differences in the occurrence of predators. However, this seems rather unlikely because predation has been found to have a small effect on tadpole survival of B. variegata 400 401 compared to drought risks (Morand 1997). In addition, after metamorphosis, B. variegata 402 displays anti-predator strategies (described in Bajger 1980) involving numerous venom glands 403 that produce a noxious secretion (König and Bininda-Emonds 2011) as well as behavioral responses such as defensive postures (the 'unken reflex', see Toledo et al. 2011) in which 404 toads exhibit their aposematic coloration. A third possible factor could be local differences in 405 the presence of competitors. However, B. variegata breeds in highly temporary water bodies 406 that are avoided by most other amphibians (Morand and Joly 1995, Morand 1997), so local 407 408 differences in the amphibian community are not likely to have an impact on the demographic 409 rate of *B. variegata*.

410

411 Contrasting patterns of environmental fluctuation influence female breeding propensity and
412 offspring productivity

413

Our results indicate that female breeding propensity was on average similar in the two
environments, but was more variable from year to year in unpredictable environments. In *B. variegata*, the decision to reproduce depends on a variety of external factors (e.g. water
deficit) and the physiological state of the female during the current breeding opportunity
(Cayuela et al. 2014), in line with the concept of 'state-based life history' (McNamara and

Houston 1996). In females, the energetic costs of breeding are usually related to three factors: 419 the storage of lipids in body fat ('capital breeder tactic'; Fitzpatrick 1976, Bonnet et al. 1998), 420 421 internal sexual mechanisms (especially oogenesis dynamics), and escaping unwanted sexual attention and possible injury inflicted by males during attempts at amplexus (Wells 2010). 422 These costs are likely to be enhanced by spatio-temporal unpredictability of breeding sites. In 423 managed woodlands in particular, breeding sites are created by logging operations that leave 424 deep ruts that fill with rainwater (Cayuela et al. 2015a) and typically remain attractive for 425 426 toads for two years. This type of water body creation generates strong local stochasticity of 427 available breeding sites. As a result, the inter-annual fidelity of females to breeding sites is 428 lower and more variable in unpredictable than in predictable environments (Pichenot 2008, Cayuela et al. *in prep*). Breeding propensity that is more variable over time in unpredictable 429 environments could thus be a result of the costs induced by moving from an aging rut network 430 to a more recent one. These costs could be indirect, due to lost metabolic energy in 431 movement, or direct, due to the time invested in movement (Bonte et al. 2012) that may 432 433 prevent a female from breeding during a reproduction opportunity. Differences in annual offspring productivity between the two types of environment 434 435 indicate contrasting lifetime schedules of resource allocation to breeding. Risk-spreading tactics, including bet-hedging and adaptive coin-flipping (see Philippi and Seger 1989, Wilbur 436 and Rudolf 2006), have been highlighted in various organisms (Veening et al. 2008, Childs et 437 438 al. 2010, Nevoux et al. 2010, Gamelon et al. 2013). In B. variegata, risk-spreading tactics 439 (e.g. egg clutch fractionation over space and time) usually allow females to cope with 440 reproductive success uncertainty caused by high desiccation risk in oviposition sites 441 (Buschmann 2002). In predictable environments, our results indicate that the inter-annual predictability of breeding site availability allows females to more consistently spread their 442 breeding effort over multiple reproduction opportunities in a year. Since female breeding 443

investment in a given reproduction period depends on past breeding status (Cayuela et al. 444 445 2014), subsequent fitness gains (number of offspring) are therefore annually lower, but more 446 constant over time in predictable environments. In contrast, in unpredictable environments, the inter-annual unpredictability of breeding site availability and localization constrains the 447 448 allocation of breeding investment over time. Since egg resorption allows females to recover energetic resources that can be reinvested later, yearly fitness gains are higher on average, but 449 less constant from year to year. Our results show that contrasting patterns of environmental 450 451 fluctuation could regulate risk-spreading tactics among populations by shaping the trade-off 452 between mean and variance of breeding investment and subsequent fitness gains.

453

454 Contrasting patterns of environmental fluctuation influence survival across ontogenetic
455 stages

456

Our results show that mean survival increases with age while temporal variation decreases, 457 458 which is in line with the environmental canalization paradigm of 'slow species' (Yoccoz and Gaillard 2003). The survival estimates of adults in our study are congruent with those found 459 460 in previous studies (= 0.80 in Plytycz and Bigaj 1993, = 0.74 in Cayuela et al. 2014). In amphibians, post-metamorphosis survival is positively correlated with age and size (Altwegg 461 and Rever 2003, Garner et al. 2011, Schmidt et al. 2012). Compared to most anurans, 462 463 recorded *B. variegata* survival rates and longevity (Smirina 1994) are high when corrected for 464 allometry (see the HAGR website: http://genomics.senescence.info). This is likely due to anti-465 predator and anti-microbial strategies – specific and unique families of antimicrobial peptides 466 in bombinatorid skin secretions protect them against invasion by pathogenic microorganisms and also play a role in protecting the animal from ingestion by predators (Conlon 2011, König 467 468 and Bininda-Emonds 2011).

At each ontogenetic stage, survival was slightly lower and more variable in 469 unpredictable environments. In these environments, the inter-annual unpredictability of 470 471 breeding sites may result in high dispersal propensity, thus reducing apparent survival (in 472 amphibians, see Perret et al. 2003, Schmidt et al. 2007); in capture-recapture approaches, apparent survival estimates include both death and permanent emigration (Nichols et al. 1994, 473 474 Kendall et al. 1997). The lower survival rate may also be the result of dispersal costs in unpredictable environments. For example, increased mortality may result from logging if 475 476 operations are carried out in the spring. Logging activities near breeding ponds may also negatively impact survival due to changes in the surface temperature of the soil and loss of 477 478 moisture (Semlitsch et al. 2009). Moreover, rut filling resulting from path restoration during reproduction activity could lead to catastrophic mortality in breeding adults. Finally, in 479 agricultural areas, prolonged exposure to pesticides can alter growth before and after 480 481 metamorphosis by inhibiting normal thyroid activity and/or increase susceptibility to disease through the suppression of the humoral response (Mann et al. 2009). 482

483

484 Conclusion

485

Taken together, the differences recorded in *B. variegata* vital rates in unpredictable and 486 487 predictable environments demonstrate that different regimes of stochastic variation in 488 environmental conditions may contribute to intraspecific variation of life history tactics along 489 the 'slow-fast' continuum. An increase in offspring productivity as well as increased variation 490 in adult survival (leading to a decrease in average annual adult survival) results in an 491 acceleration of the life history cycle in populations in unpredictable artificial environments. 492 Habitat anthropization modifies the spatio-temporal availability of breeding sites and contributes to divergent selection gradients at intra-specific scales. Because of this new 493

selective pressure, *B. variegata* colonization of human-impacted habitats may lead to local
adaptations that could minimize to some extent the fitness costs induced by changes in the
regime of environmental fluctuation. Understanding the processes of this adaptation
(plasticity vs. selection) requires further investigation based on experimental and/or
translocation approaches.

499

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501

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SUPPLEMENTAL MATERIAL

Appendix A

724 Survey design and Environmental characteristics of the five studied populations of *Bombina*

725 variegata.

- 726 Appendix B
- Assessing a potential bias of vital rates resulting from the violation of the population closure

728 assumption

- TABLE 1. Survey design characteristics for the five populations investigated in the study:
- study period, survey duration, number of capture sessions performed during the survey period,
- total number of captures, and number of individuals identified during the survey.

	Study	Survey	Capture	Number of	Number of
	period	duration	sessions	captures	individuals
POP1	2000–2008	8 years	45	953	581
POP2	2008–2014	7 years	23	12300	9418
POP3	2005–2014	10 years	111	1053	456
POP4	2010-2014	5 years	40	3053	1154
POP5	2010-2014	5 years	37	2086	758

733

TABLE 2. Definition of parameters estimated in the capture–recapture multi-event model.

Parameter	Definition
α	probability that a newly metamorphosed individual is recruited (FIG.1, matrix A)
β	probability that an individual does not reach the next ontogenetic stage (FIG.1, matrix A)
φ	probability that an individual survives (FIG.1, matrix Φ)
γ^B	probability that a sexually individual bred at time t, given it has bred at $t - 1$ (FIG.1,
	matrix Y)
γ^{NB}	probability that a sexually mature individual skipped a breeding opportunity at time t,
	given it skipped breeding at $t - 1$ (FIG.1, matrix Y)
p	probability that an individual was recaptured during a secondary session (FIG.1, matrix P)
	preprint

TABLE 3. Model selection procedure showing the competing models run for each population: r = model rank, k = number of parameters, Dev. = residual deviance, AICc = Akaikeinformation criteria adjusted for a small sample size and their derivate quantities, w = AICcrelative weight.

r		Mode	l definition		k	Dev.	AICc	W						
	Survival	Breeding	Recruitment	Recapture	C									
PC	OP1 (Unpredictabl	e)												
1	$AGE \times YEAR$	YEAR	YEAR	$AGE \times SEX + YEAR$	59	11,942.87	12,064.39	1.00						
2	$AGE \times YEAR$	SEX	YEAR	$AGE \times SEX + YEAR$	50	11,973.54	12,076.38	0.00						
3	$AGE \times YEAR$	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	75	11,929.29	12,085.74	0.00						
4	$AGE \times YEAR$	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	YEAR	72	11,954.25	12,104.19	0.00						
5	YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	64	12,006.38	12,139.06	0.00						
6	AGE	$SEX \times YEAR$	YEAR	$AGE \times SEX + YEAR$	61	12,100.12	12,226.37	0.00						
7	AGE × YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	AGE × SEX	67	12,556.51	12,705.64	0.00						
PC	OP2 (Unpredictabl	e)												
1	AGE × YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	61	127,315.07	127,437.49	0.92						
2	AGE × YEAR	YEAR	YEAR	$AGE \times SEX + YEAR$	49	127,331.91	127,430.19	0.08						
3	$AGE \times YEAR$	SEX	YEAR	$AGE \times SEX + YEAR$	44	127,364.50	127,452.73	0.00						
4	AGE × YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	YEAR	58	127,603.53	127,719.91	0.00						
5	YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	50	128,088.70	128,188.98	0.00						
6	AGE	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	47	128,126.73	128,220.98	0.00						
7	$AGE \times YEAR$	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX$	55	128,228.80	128,339.15	0.00						
PC	OP3 (Unpredictabl	e)												
1	$AGE \times YEAR$	YEAR	YEAR	$AGE \times SEX + YEAR$	67	19,925.58	20,062.09	0.58						
2	$AGE \times YEAR$	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	85	19,888.94	20,062.98	0.42						
3	$AGE \times YEAR$	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	YEAR	83	19,908.03	20,077.88	0.00						
4	YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	70	19,946.22	20,088.96	0.00						
5	AGE	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	64	19,986.76	20,117.05	0.00						
6	$AGE \times YEAR$	SEX	YEAR	$AGE \times SEX + YEAR$	56	20,010.99	20,124.74	0.00						
7	$AGE \times YEAR$	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX$	76	20,730.38	20,885.61	0.00						
PC	OP4 (Predictable)													

1	$AGE \times YEAR$	SEX	YEAR	$AGE \times SEX + YEAR$	34	40,734.24	40,802.44	1.00
2	AGE	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	35	41,777.84	41,848.05	0.00
3	AGE × YEAR	YEAR	YEAR	$AGE \times SEX + YEAR$	35	42,156.09	42,226.3	0.00
4	AGE × YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	43	42,149.91	42,236.23	0.00
5	YEAR	$SEX \times YEAR$	YEAR	$AGE \times SEX + YEAR$	36	42,183.84	42,256.06	0.00
6	AGE × YEAR	$SEX \times YEAR$	YEAR	$AGE \times SEX$	39	42,323.67	42,401.93	0.00
7	AGE × YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	YEAR	40	42,412.86	42,493.13	0.00
PO	P5 (Predictable)							
1	$AGE \times YEAR$	SEX	YEAR	$AGE \times SEX + YEAR$	34	25,874.07	25,942.37	1.00
2	AGE	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	34	26,963.43	27,031.73	0.00
3	AGE × YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	43	26,995.75	27,082.23	0.00
4	$AGE \times YEAR$	YEAR	YEAR	$AGE \times SEX + YEAR$	35	27,019.16	27,089.48	0.00
5	YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	$AGE \times SEX + YEAR$	36	27,049.63	27.121.97	0.00
6	AGE × YEAR	$\mathbf{SEX} \times \mathbf{YEAR}$	YEAR	YEAR	49	27,199.63	27,298.25	0.00
7	AGE × YEAR	$SEX \times YEAR$	YEAR	$AGE \times SEX$	38	27,255.67	27,332.05	0.00

- FIG.1. Elementary matrices for state–state and events transitions: UM = unborn male, UF =
- unborn female, JM = juvenile male, JF = juvenile female, SM = subadult male, SF = subadult
- female, BM = breeding male, BF = breeding female, NBM = non-breeding male, NBF = non-
- breeding female, $\dagger = \text{dead}$, NS = not seen, J = seen juvenile, S = seen subadult, M = seen
- sexually mature male, F = seen sexually mature female.
- FIG. 2. Conditional breeding probability (breeding or skipping breeding during two
- consecutive years) and temporal variation ratio of breeding propensity of *B. variegata* females
- studied in France over the 2000–2014 period. The 95% CI is shown using error bars.
- FIG. 3. Annual offspring productivity (number of recruited females per female per year) and
- temporal variation (standard deviation) in five populations of *B. variegata* studied in France
- 753 over the 2000–2014 period. The 95% CI is shown using error bars.
- FIG. 4. Mean annual survival at three ontogenetic stages (juvenile, subadult and adult) and
- temporal variation ratios in five populations of B. variegata studied in France over the 2000-
- 756 2014 period. The 95% CI is shown using error bars.

matrix A (recruitment and age transition)

	UM	$U\!F$	JM	JF	SM	SF	BM	BF 1	VBM	N	BF ·	Ť	UM	UF	JM	JF	SM	SF	BM	BF	NBM	NBF	+
UM	$/1 - \alpha^{JM}$	0	α^{JM}	0	0	0	0	0	0	0	0\	UM	/1	0	0	0	0	0	0	0	0	0	0 \
UF		$1 - \alpha^{JF}$	0	α^{JF}	0	0 0	Ő	0	Õ	0	$\tilde{0}$	UF	0	1	0	0	0	0	0	0	0	0	0
JM	0	0	β^{JM}	0	$1 - \beta^{JM}$	0	0	0	0	0	0	JM	0	0	Φ^{JM}	0	0	0	0	0	0	0	$1 - \phi^{JM}$
IF	0	0	0	ß ^{JF}		$1 - \beta^{JF}$	0	0	0	0	0	JF	0	0	0	Φ^{JF}	0	0	0	0	0	0	$1 - \phi^{JF}$
SM		0	0	Р 0	в SM	1 p 0	$1 - \beta^{SM}$	0	0	0	0	SM	0	0	0	0	ϕ^{SM}	0	0	0	0	0	$1 - \phi^{SM}$
SIVI		0	0	0	Р 0	ßSF	1 p 0	$1 - R^{SF}$	0	0	0	SF	0	0	0	0	0	Φ^{SF}	0	0	0	0	$1 - \phi^{SF}$
BM		0	0	0	0	р 0	1	тр 0	0	0	0	BM	0	0	0	0	0	0	Φ^{BM}	0	0	0	$1 - \phi^{BM}$
BF	0	0	0	0	0	0	0	1	0	0	0	BF	0	0	0	0	0	0	0	Φ^{BF}	0	0	$1 - \phi^{BF}$
NBM	0	0	0	0	0	0	0	0	1	0	0	NBM	0	0	0	0	0	0	0	0	Φ^{NBM}	0	$1 - \phi^{NBM}$
NBF	0	0	0	0	0	0	0	0	0	1	0	NBF	0	0	0	0	0	0	0	0	. 0	Φ^{NBF}	$1 - \Phi^{NBF}$
Ť	\ 0	0	0	0	0	0	0	0	0	0	1/	ţ	$\setminus 0$	0	0	0	0	0	0	0	0	0	i /

matrix Y (breeder/non-breeder transition)

	UM	UF	JM	JF	SM	SF	BM	BF	NBM	NBF	Ť
$U\!M$	/1	0	0	0	0	0	0	0	0	0	0
UF	0	1	0	0	0	0	0	0	0	0	0
JM	0	0	1	0	0	0	0	0	0	0	0
JF	0	0	0	1	0	0	0	0	0	0	0
SM	0	0	0	0	1	0	0	0	0	0	0
SF	0	0	0	0	0	1	0	0	0	0	0
BM	0	0	0	0	0	0	γ^{BM}	0	$1 - \gamma^{BM}$	0	0
BF	0	0	0	0	0	0	0	γ^{BF}	0	$1 - \gamma^{BF}$	0
NBM	0	0	0	0	0	0	$1 - \gamma^{NBM}$	0	γ^{NBM}	0	0
NBF	0	0	0	0	0	0	0	$1 - \gamma^{NBF}$	0	γ^{NBF}	0
+	/0	0	0	0	0	0	0	0	0	0	1/

matrix Φ (survival transition)

matrix P (recapture)

	NS	J	S	M	F
UM	/ 1	0	0	0	0 \
$U\!F$	1	0	0	0	0
JM	$1 - p^{JM}$	p^{JM}	0	0	0
JF	$1 - p^{JF}$	p^{JF}	0	0	0
SM	$1 - p^{SM}$	0	p^{SM}	0	0
SF	$1 - p^{SF}$	0	p^{SF}	0	0
BM	$1 - p^{BM}$	0	0	p^{BM}	0
BF	$1 - p^{BF}$	0	0	0	p^{BF}
NBM	1	0	0	0	0
NBF	1	0	0	0	0 /
Ť	\ 1	0	0	0	0 /





